THE FACTORIAL STRUCTURE OF JOB-RELATED AFFECTIVE WELL-BEING: POLISH ADAPTATION OF THE WARR’S MEASURE

EMILIA MIELNICKU and MARIOLA ŁAGUNA
The John Paul II Catholic University of Lublin, Lublin, Poland
Institute of Psychology, Department of General Psychology

Abstract
Objectives: The first aim of the study reported in this article was to test the factorial structure of job-related affect in a Polish sample. The second aim was to develop the Polish adaptation of the Warr’s job-related affective well-being measure published in 1990, which is designed to assess 4 types of affect at work: anxiety, comfort, depression, enthusiasm. Material and Methods: A longitudinal study design with 2 measurement times was used for verifying the psychometric properties of the Polish version of the measure. The final sample consisted of 254 Polish employees from different professions. Participants were asked to fill in a set of questionnaires consisting of measures capturing job-related affective well-being, mood, and turnover intention. Results: The first step of analysis was to test the theoretically-based structure of the job-related affective well-being measure in a Polish sample. The confirmatory factor analysis revealed that a 4-factor model best describes the structure of the measure in comparison to 5 alternative models. Next, reliability of this measure was assessed. All scales achieved good internal consistency and acceptable test-retest reliability after 2 weeks. Finally, the convergent and discriminant validity as well as the criterion and predictive validity of all job-related affective well-being scales was confirmed, based on correlations between job-related affect and mood as well as turnover intention. Conclusions: The results suggest that the Polish adaptation of Warr’s job-related affective well-being measure can be used by scientists as well as by practitioners who aim at assessing 4 types of affective well-being at a work context. Int J Occup Med Environ Health 2018;31(4):429–443

Key words: Mood, Employees, Turnover intention, Circumplex model of affect, Job-related affective well-being, Longitudinal research

INTRODUCTION
For a long time, there was a dominant opinion that emotions and rational thinking and acting are mutually exclusive. That is why researchers did not perceive emotions as important factors that may be related to job performance [1]. Nowadays, however, the interest in affect in the organizational context is increasing, since there is growing evidence that emotional reactions are connected with rational decision making [2], as well as with health [3], and different work outcomes [4]. For these reasons the interest in tools dedicated to the assessment of affect at work is rising, especially that there are only a few measures capturing affect in a work context available [5,6]. A broadly used method [4,7,8] is the job-related affective well-being measure developed by Warr [9]. The aim of this article is to analyze the factorial structure and psychometric properties of this tool using longitudinal data from Polish employees. We present the Polish adaptation of Warr’s
measure of job-related affective well-being dedicated to capturing 4 dimensions of affect at work.

**Affect and its role in a work context**

Dictionary definitions of “emotions,” “affect,” and “feelings” overlap and there is no clear distinction between them [e.g., 10]. Similarly, researchers propose diverse definitions distinguishing (or not) these concepts [11]. In this paper we follow the conceptualization of affect proposed by Fredrickson [12]. According to her theory, when something important happens, it triggers numerous affective responses. Emotions are relatively short in duration, and they may be conscious or unconscious. Affect is a more general concept than emotions and moods, and it refers to consciously accessible feelings. Unlike emotions, which always have an object, affect may be objectless [12]. Following this conceptualization, we treat affect as consciously accessible feelings, evident in moods and emotions.

Research findings increasingly show that affect has an important influence on individuals’ behavior, including a work context [3,13,14]. According to the broad theory, positive emotions predict positive outcomes by broadening the way individuals process information and increasing the number of goals they want to pursue [15]. The theory holds that such a broadening process, which is triggered by positive emotions, helps to build more durable personal resources, including physical, social, intellectual, and psychological ones [12]. In the occupational context, it has been revealed that positive emotions influence a variety of performance-relevant outcomes such as judgments, creativity, helping behavior, and risk taking [16], help to reduce occupational stress [14] and are connected with work engagement [4,17]. Negative emotions also play an important role in an organizational context. They are strongly connected with occupational stress and may lead to psychological breakdown [3] and burnout [18]. What is more, they relate to workplace incivility [19] and various counterproductive work behaviors, such as mobbing, antisocial behavior, aggression or withdrawal [20]. There is also evidence that positive affect towards personal work-related goals is positively related whereas negative affect is negatively related to the goal achievement [21]. All those research findings clearly demonstrate that it is important to take affect into account when investigating activity at work.

**Conceptualizations of affective dimensions**

The considerable interest in subjective dimensions has led to several conceptualizations of affect and its dimensions. Researchers have proposed diverse theories concerning the structure of affect and suggestions about how affect should be captured. For instance, Diener et al. [22] claim that positive and negative affect should be treated as 2 poles of 1 dimension since these 2 states are unlikely to be experienced at the same time. Their study has revealed that the more frequently positive affect occurs, the less frequently negative affect is experienced.

This one-dimensional approach is, however, criticized by other researchers who claim that positive and negative affect are 2 independent constructs [23]. Warr et al. [24] claim that the correlation between numbers of desirable and undesirable episodes of life events as well as experienced feelings is not high enough to make a prediction about duration of negative affect based on duration of positive affect. According to Argyle and Martin [25], positive and negative affect have different causes, and Davidson [26] has revealed that positive affect is connected with left frontal activation of the brain, while negative emotions are related to right frontal activation.

Nowadays, a more popular notion is that diverse affective states are related to each other in a highly systematic way [27]. They are often categorized along 2 dimensions: pleasantness and emotional activation [28]. Pleasantness divides emotions into positive and negative ones whereas activation refers to high or low arousal. These are basic dimensions of the widely acknowledged and empirically supported circumplex model of affect [27,29]. This cir-
Warr’s job-related affective well-being measure

Based on this model, Warr [9] created the measure of job-related affective well-being intended to capture 4 affective dimensions: anxiety, comfort (labelled earlier contentment), depression, and enthusiasm. These scales capture the endpoints of 2 circumplex model axes: anxiety-comfort and depression-enthusiasm, while a third axis: displeased-pleased is considered to depict job satisfaction which is measured by other scales [9]. The Warr job-related affective well-being measure is a short instrument built of 12 items, each of them being a single adjective [9]. The respondents evaluate on a six-point scale (1 – never, 6 – all of time) how frequently their work made them feel certain feelings during the past few weeks. There are 6 positive (calm, contented, relaxed, cheerful, enthusiastic, optimistic) and 6 negative (tense, uneasy, worried, depressed, gloomy, miserable) feelings. The score on each of 4 proposed scales of the instrument is a mean or a sum of 3 items. The time allotted for filling in the measure is very short and does not exceed 5 min, which makes the instrument very useful in research with many measures and in on-line questionnaires. As this measure has clear theoretical underpinnings it has become very popular in research in work and organizational psychology [4,7,8].

There are also other measures of affect available. One of the most popular, the Positive and Negative Affect Schedule (PANAS) [6], captures positive and negative general, context-free affect. The University of Wales Institute of Science and Technology (UWIST) Mood Adjective Checklist (UMACL) [30], which consists of 3 dimensions – energetic arousal, tense arousal and hedonic tone – is designed to measure rather short-lived feeling states like mood. However, none of these scales is designed specifically for measuring affect in a work context. They also do not cover the whole spectrum of affect depicted by the circumplex model (e.g., PANAS does not cover low activation affect).

The measure which captures affect in a work context is the Job-related Affective Well-being Scale (JAWS) [5]. It also

Cumplex model was applied by Warr [9] to the work context. In his conception job-related affective well-being is defined by 2 principal dimensions of pleasure (horizontal) and arousal (vertical) (Figure 1), however, according to Warr [9], arousal by itself is not a reflection of affect. This model allow to describe both the content and the intensity of job-related affective states and to represent affective well-being along 3 key axes: displeased-pleased, anxiety-contentment (called also anxiety-comfort), and depression-enthusiasm. Thanks to this it is possible to describe not only valence, i.e., positive and negative affect but also to distinguish diverse content of affective experiences at work. In consequence each specific job-related affect can be understood as a linear combination of both valence and arousal, and arise from “cognitive interpretations of core neural sensations that are the products of 2 independent neurophysiological systems” [29, p. 715].


HANA – activated negative affect; HAPA – activated positive affect; LANA – low activation negative affect; LAPA – low-activation positive affect.

Fig. 1. Four-factor model of Warr’s job-related affective well-being measure [9]
refers to Warr’s model of affective well-being at work [5] and assess people’s emotional reactions to their job in the past month which are categorized along 2 dimension: pleasurableness (positive and negative affect) and arousal (high and low intensity). This hypothesized 4-factor structure was confirmed in a Polish sample of police officers [31]. As this method is longer (30 items in a full and 20 items in a short version) than the Warr [9] job-related affective well-being measure, it is less popular in research. What is more, despite the fact that the original version has good reliability [5], there are same psychometric problems in various non-English versions [31]. The advantage of the Warr measure is also that it allows to capture diverse content of 4 affective experiences at work which have clear theoretical foundations in the circumplex model of affect [27,29]. Having already the Polish adaptation of the JAWS scale, we have developed the Polish adaptation of Warr’s instrument. As a Polish name of the Warr measure we propose to use Kwestionariusz Afektu w Pracy (KAP). This name allows us to distinguish the measure from JAWS, which is called in Polish Skala Dobrostanu Emocjonalnego w Pracy [32].

The popularity of the Warr measure raises the issue of how many and what dimensions should be used for describing subjective-well-being in the best way. A few alternative models with the different number of factors that may be captured by Warr’s instrument were tested. In previous analyses of this measure, few alternative models were tested (Table 1), using the exploratory and confirmatory factor analysis (CFA). Daniels et al. [33] revealed a 3-factor structure (positive affect, negative affect, and pleasure-displeasure factor) as the best fitting to the data. Others suggest to treat positive and negative affect as 2 main dimensions of well-being [34]. Another factor model with 2 dimensions (anxiety-comfort and depression-enthusiasm) proposed by Warr [9] has also received support in several studies [9,33,35,36]. Other studies have revealed only 1 common latent factor [8]. Two recent studies [7,37] using the CFA, identified the 4-factor model as the best representation of the structure of affective well-being in a work context.

In this study, 6 alternative factor models of job-related affective well-being were tested in a Polish sample: the 4-factor model (anxiety, comfort, depression, enthusiasm), based on Warr’s theory, a 2-factor model (dimensions anxiety-comfort, depression-enthusiasm), another alternative 2-factor model (positive affect, negative affect), a 3-factor model (positive affect, negative affect, and pleasantness-unpleasantness), and finally a 1-factor model. On the basis of the previous research [7,37] and Warr’s [9] theory we hypothesize that a 4-factor solution will best fit the data.

Since its publication in 1990, the Warr measure has been used in numerous studies and translated into many languages, for instance Finnish [7], Portuguese [38], and Spanish [35]. What is more, it is widely used in studies of various occupational groups (Table 1) such as managers and professional workers, blue-collar and white-collar employees, people working in education, and police officers [7,35–37,39]. Development of the Polish adaptation of the measure will allow comparisons of results from Polish samples with results from other studies. Therefore in the subsequent sections we present a study attesting to the psychometric properties of the Polish adaptation of Warr’s [9] job-related affective well-being measure.

**Associates of affective well-being**

To assess criterion and predictive validity of the Polish version of the Warr measure, we investigated correlations of its scales with 2 constructs, namely mood and turnover intention. The first one, positive and negative mood, may be captured as a context-free phenomenon, having, however, close relationships with job-related affect, as we show below. The other, turnover intention, is strictly related to a work context.

Mood may be defined as a short-term affect and it is a quick response to an environmental stimulus [40]. A-
### Table 1. Summary of previous studies on the Warr's job-related affective well-being measure [9]

<table>
<thead>
<tr>
<th>Model tested and other characteristics of studies</th>
<th>Exploratory factor analysis studies</th>
<th>Confirmatory factor analysis studies</th>
</tr>
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<tbody>
<tr>
<td>Model 1. Four correlated factors: anxiety, comfort, depression, enthusiasm</td>
<td></td>
<td>X</td>
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<tr>
<td>Model 2. Two correlated factors: anxiety-comfort, depression-enthusiasm</td>
<td>X</td>
<td>X, B</td>
</tr>
<tr>
<td>Model 3. Two correlated factors: positive affect, negative affect</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Model 4. Three factors: positive affect, negative affect, pleasantness-unpleasantness</td>
<td></td>
<td></td>
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<tr>
<td>Model 5. Single factor: well-being</td>
<td></td>
<td>X, A</td>
</tr>
<tr>
<td>Model 6. Hierarchical: Four factors and higher order factor well-being</td>
<td></td>
<td></td>
</tr>
</tbody>
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<tr>
<th>Sample</th>
<th>1 686 employees from different professions</th>
<th>299 Spanish factory workers and blue-collar employees</th>
<th>3 044 civil servants and white-collar workers</th>
<th>267 UK civil servants, 156 UK police officers</th>
<th>615 Finnish managers</th>
<th>1 446 Portuguese police officers</th>
</tr>
</thead>
<tbody>
<tr>
<td>Reliability of the best fitting model scales</td>
<td>anxiety-comfort: 0.76, depression-enthusiasm: 0.80</td>
<td>anxiety-comfort: 0.86, depression-enthusiasm: 0.77</td>
<td>anxiety-comfort: 0.82, depression-enthusiasm: 0.85</td>
<td>not reported</td>
<td>anxiety: 0.74/0.78, comfort: 0.80/0.83, depression: 0.83/0.85, enthusiasm: 0.82/0.85</td>
<td>anxiety: 0.87, comfort: 0.88, depression: 0.91, enthusiasm: 0.92</td>
</tr>
</tbody>
</table>

X – model tested; B – model with the best fit; A – model with acceptable fit.

* Reliability: Time 1/Time 2.
ccording to the theory, work and non-work events have an impact on employees’ affect both immediately and over a longer period of time [41]. This, in turn, influences work attitudes, performance and other kinds of affect-driven behavior [42]. Miner et al. [43] have revealed that mood largely determines the intensity of employees’ reactions to events, depending on how good someone’s mood is, he or she may experience work events differently. Rothbard and Wilk [40] have shown that start-of-workday mood of employees in call centers may predict their affect on subsequent calls and their perception of customers’ affective display. Start-of-workday positive mood was positively related to employees’ positive affect before calls as well as their perceptions of customers’ positive affective display. In contrast, if employees started their workday with a negative mood they felt negative affect on subsequent calls and they claimed that customers also experienced a bad mood. Consistent with this, Iles and Judge [44] observed individual variations in job satisfaction across time, as their within-individual analyses showed that individuals’ job satisfaction varied in synchrony with their mood. Based on these findings, mood appears to act as an “affective prime” [43]. That is why in the current research we assume that employees’ current mood is related to and may influence how they assess their affective well-being at work. More precisely, we expect that positive mood will be accompanied by enthusiasm and comfort dimensions while negative mood will be accompanied by anxiety and depression dimensions as captured by the Warr measure.

Turnover is a major concern for most organizations. It is perceived as an expression of withdrawal behavior in the workplace next to the absenteeism, lateness or tardiness [20]. Turnover intention is an employee’s desire to leave an organization [45]. It is a one-step before quitting from the organization, as it has been acknowledged as the best predictor of actual turnover [46,47]. The relationship between affect and turnover has been widely researched and generally the results show that positive emotions lead to positive outcomes in a work context and negative emotions bring unwanted effects [3,4,14,18]. More specifically, they showed that positive affect lowers turnover intention, while negative affect fosters it [20,45,48]. Based on the previous research we expect that there will be positive relationship between anxiety and depression and turnover intention, and negative relationship between enthusiasm and comfort and turnover intention.

MATERIAL AND METHODS

Procedure and participants

A longitudinal study with 2 measurement times at 2 week intervals was carried out. Data was collected anonymously, and participants provided their personal codes which enabled us to match their data from the first and second study waves. Involvement in the study was voluntary and there was no reward for participants. The researchers visited participants in their workplace or during their postgraduate trainings and asked them if they would be willing to complete a questionnaire.

In the first study wave (Time 1) 380 Polish employees from various professions were asked to complete the paper-and-pencil questionnaires. Two weeks later (Time 2) questionnaires were given to all the respondents who had participated in the first study wave. The response rate at time 2 was 67.11% (N = 255). Most of the analyses presented in this paper are based on the data from the first study wave. Univariate analyses of variance indicated that there were no significant differences in gender, age, and main work characteristics between the participants who took part in both study waves and participants who dropped out after Time 1 (all p > 0.05).

The whole sample in Time 1 comprised 133 men (35%) and 247 women (65%). The age of participants ranged from 18 to 64 years old, average age was 32.81 years old (standard deviation (SD) = 8.90). In the total sample 54.7% had full-time work contract, 23.2% had part-time work and 22.1% had other types of job agreement.
Most of the employees (91.3%) worked in the service sector, 5.5% in the construction sector and 3.2% in industry.

**Measures**

Job-related affect was measured by Warr’s job-related affective well-being measure described in detail in the Introduction section. The instrument was translated into Polish by 4 independent translators. Based on these translations, the Polish version was elaborated, and then back-translated. The final Polish version of the items is presented in the Table 2. The scales’ reliability is presented in the “Results” section.

Mood was evaluated using a 10 item scale developed by Wojciszke and Baryła [49]. Participants rated their agreement with statements describing their mood. The responses to each item were given on a five-point scale (from 1 – “I do not agree” to 5 – “I agree”). There were 5 statements covering positive mood (e.g., “I feel excellent”) and 5 describing negative mood (e.g., “I am in a bad mood”). A higher score indicates a higher level of each kind of mood. Cronbach’s α is 0.92 for negative mood and 0.91 for positive mood at Time 1 and 0.93, and 0.92 respectively at Time 2, confirming high reliability of both scales.

Turnover intention was measured by a three-item measure [50]. An example item is “I think a lot about leaving the organization.” Each statement was responded to on a 5-point scale (ranging from 1 – “strongly disagree” to 5 – “strongly agree”). A higher score indicates a higher intention to leave an organization. Cronbach’s α is 0.90 at

| Table 2. Descriptive statistics and correlations between the items of Warr’s job-related affective well-being measure [9] in a sample of Polish employees (N = 380) |
|---------------------------------|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|
| Original item and its Polish version* | M   | SD  | 1   | 2   | 3   | 4   | 5   | 6   | 7   | 8   | 9   | 10  |
| 1. Relaxed / Zrelaksowany | 3.15 | 1.21 | 1   |     |     |     |     |     |     |     |     |     |
| 2. Worried / Zmartwiony | 2.55 | 1.10 | -0.38 | 1   |     |     |     |     |     |     |     |     |
| 3. Depressed / Przygnębiony | 2.22 | 1.13 | -0.36 | 0.61 | 1   |     |     |     |     |     |     |     |
| 4. Calm / Spokojny | 3.75 | 1.04 | 0.66 | -0.50 | -0.48 | 1   |     |     |     |     |     |     |
| 5. Content / Zadowolony | 3.78 | 1.13 | 0.68 | -0.41 | -0.49 | 0.73 | 1   |     |     |     |     |     |
| 6. Gloomy / Posępny | 2.91 | 1.08 | -0.35 | 0.56 | 0.82 | -0.48 | -0.51 | 1   |     |     |     |     |
| 7. Optimistic / Optymistyczny | 3.81 | 1.27 | 0.62 | -0.36 | -0.43 | 0.62 | 0.74 | -0.44 | 1   |     |     |     |
| 8. Tense / Spięty | 2.83 | 1.10 | -0.45 | 0.62 | 0.48 | -0.48 | -0.38 | 0.45 | -0.32 | 1   |     |     |
| 9. Enthusiastic / Entuzjastyczny | 3.58 | 1.29 | 0.64 | -0.31 | -0.38 | 0.55 | 0.73 | -0.42 | 0.84 | -0.29 | 1   |     |
| 10. Cheerful / Radosny | 3.77 | 1.16 | 0.61 | -0.34 | -0.42 | 0.62 | 0.75 | -0.45 | 0.79 | -0.34 | 0.84 | 1   |
| 11. Miserable / Nieszczęśliwy | 1.79 | 1.09 | -0.33 | 0.58 | 0.74 | -0.46 | -0.50 | 0.74 | -0.46 | 0.44 | -0.41 | -0.42 | 1   |
| 12. Uneasy / Niespokojny | 2.77 | 1.12 | -0.44 | 0.71 | 0.57 | -0.53 | -0.42 | 0.53 | -0.39 | 0.79 | -0.33 | -0.39 | 0.51 | 1   |

M – mean; SD – standard deviation.
** All correlations are statistically significant at the level of at least p < 0.01.
Data analysis strategy
First, descriptive statistics and correlations between single items were analyzed. Next, to examine the factorial structure of the job-related affective well-being measure – the CFA was conducted using AMOS [51]. Finally, to assess psychometric properties of scales emerging from the best fitting CFA model, the descriptive statistics, internal consistency, temporal stability and validity of the scales were analyzed. These analyses were performed by using IBM SPSS 23.

RESULTS
Descriptive statistics and correlations between items
The initial analysis concerning items in the Polish version of the Warr [9] measure was performed on the data from the first study wave (Time 1). The mean score of items ranged from 1.79 for item 9 to 3.78 for item 5 on a six-point scale showing no extremely low or high mean values (Table 2). As it can be expected, correlations between items describing positive and negative affect are negative, ranged from –0.29 to –0.5, and the correlations between items with the same valence are positive, ranging from 0.41 to 0.84, all being statistically significant.

Comparison of alternative factor models
To test the factorial structure of the job-related affective well-being measure in a Polish sample of employees, 6 alternative CFA models were tested. The first model (Model 1) consisted of 4 correlated latent factors of anxiety, comfort, depression and enthusiasm. The second model (Model 2) contained 2 correlated latent factors of anxiety-comfort and depression-enthusiasm. The third model (Model 3) also comprised 2 correlated latent factors labeled positive affect and negative affect. The fourth model (Model 4) consisted of the 3 factors of positive affect, negative affect, and pleasantness-unpleasantness. The fifth model (Model 5) included one latent factor. Finally, the sixth hierarchical model (Model 6) consisted of 4 latent factors (the same as in Model 1) and a higher order latent factor named well-being.

The models were estimated using maximum likelihood estimation on the data from the first study wave. Missing data, which did not exceed 1.5% in any item, was handled using regression imputation. The model fit was assessed by using the Chi² Goodness-of-Fit Statistic, the Root Mean Square Error of Approximation (RMSEA), the Standardized Root Mean Square Residual (SRMR), the Comparative Fit Index (CFI), and the Akaike Information Criterion (AIC). An acceptable model fit was indicated by values below 0.08 for RMSEA and SRMR [52], values greater than 0.90 for the CFI [53,54] and the lower AIC index, the better is the model fit [55]. The alternative nested models were compared using the Chi²-difference test (ΔChi²) and difference in CFI (ΔCFI) [52,55]. An absolute difference in CFI that is higher than 0.01 (ΔCFI > 0.01) would indicate a significant difference in model fit [56]. We have not improved any model based on modification indices. Examination of alternative models shows that only the 4-correlated-factor model (Model 1) obtained acceptable fit in nearly all fit indices (Table 3), except for the RMSEA which slightly exceeded the criterial value of 0.08. None of the alternative models reached the criteria of acceptable fit [52,55]. What is more, tests of differences confirm that the 4 correlated factor model fits the data significantly better than all other alternative models (for all comparisons p value for ΔChi² exceeded 0.05 and ΔCFI > 0.01).

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Descriptive statistics and reliability of scales

As we had confirmed the factorial structure of Warr’s [9] instrument in a Polish sample, we next analyzed psychometric properties of its scales. Basic descriptive statistics and correlations between scales for both measurement times are presented in the Table 4. Correlations for scales ranged from –0.4 (between anxiety and enthusiasm) to 0.78 (between enthusiasm and comfort) at Time 1, and they ranged from –0.47 (between depression and enthusiasm) to 0.82 (between enthusiasm and comfort) at Time 2.

The reliability of the scales was assessed by evaluation of their internal consistency and test-retest stability. Cronbach α values (Table 4) show that all scales achieved good internal consistency as they ranged from 0.87 to 0.93 at Time 1 and from 0.86 to 0.93 at Time 2. The highest reliability was for the enthusiasm scale (in both measurement times), and the lowest indicators were obtained for the scale measuring comfort at Time 1 and anxiety at Time 2.

To assess test-retest temporal stability after 2 weeks, the Pearson correlation coefficients between 2 measurement times were calculated for each scale. Test–retest correlations were 0.76 for enthusiasm, 0.72 for comfort, 0.68 for depression, and 0.65 for anxiety, all being statistically significant. These correlation coefficients demonstrate that the most stable scores were those for the enthusiasm scale. Although the anxiety scale scores were the least stable, they still indicated acceptable temporal stability, especially taking into account that the scales are very short.

Convergent and discriminant validity

Convergent validity is achieved when 3 requirements are fulfilled: all factor loadings are statistically significant, the values of average variance extracted (AVE) for every construct exceed 0.5, composite reliability (CR) values exceed 0.6 [57]. The results confirmed the convergent validity of all scales. In the best fitting model factor loadings for anxiety...
They have been calculated separately for each study wave (Time 1 and Time 2). At both measurement times the values of correlation coefficients between enthusiasm and positive mood as well as between comfort and positive mood are statistically significant and positive (Table 5) whereas these 2 dimensions of affect are correlated negatively with negative mood. For anxiety and depression scales there are negative correlations with positive mood, and positive correlations with negative mood. Turnover intention is correlated positively with anxiety and depression dimensions and negatively with enthusiasm and comfort dimensions in both measurement times. All these correlation coefficients are statistically significant, and point out that criterion validity is reached for all scales.

Predictive validity has been assessed based on correlations between mood measured at Time 1 and job-related affect measured at Time 2. It has been assumed that positive mood will predict a higher level of enthusiasm and comfort and lower level of anxiety and depression, while negative mood has been hypothesized to show opposite

Table 4. Internal consistency, descriptive statistics and correlations between scales of the Warr’s job-related affective well-being measure [9] in a sample of Polish employees

<table>
<thead>
<tr>
<th>Study wave and job-related affect</th>
<th>α</th>
<th>M</th>
<th>SD</th>
<th>SKE</th>
<th>K</th>
<th>Correlation*</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Time 1 (N = 380)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1. Anxiety</td>
<td>0.88</td>
<td>2.71</td>
<td>0.99</td>
<td>0.55</td>
<td>0.08</td>
<td></td>
</tr>
<tr>
<td>2. Comfort</td>
<td>0.87</td>
<td>3.56</td>
<td>1.00</td>
<td>–0.05</td>
<td>–0.30</td>
<td>–0.55</td>
</tr>
<tr>
<td>3. Depression</td>
<td>0.91</td>
<td>2.01</td>
<td>1.01</td>
<td>1.21</td>
<td>1.24</td>
<td>0.64 –0.53</td>
</tr>
<tr>
<td>4. Enthusiasm</td>
<td>0.93</td>
<td>3.72</td>
<td>1.16</td>
<td>–0.26</td>
<td>–0.43</td>
<td>–0.40 0.78 –0.49</td>
</tr>
<tr>
<td>Time 2* (N = 255)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1. Anxiety</td>
<td>0.86</td>
<td>2.58</td>
<td>0.95</td>
<td>0.53</td>
<td>–0.09</td>
<td></td>
</tr>
<tr>
<td>2. Comfort</td>
<td>0.88</td>
<td>3.60</td>
<td>1.02</td>
<td>0.03</td>
<td>–0.43</td>
<td>–0.60</td>
</tr>
<tr>
<td>3. Depression</td>
<td>0.91</td>
<td>2.01</td>
<td>0.93</td>
<td>1.07</td>
<td>0.70</td>
<td>0.64 –0.50</td>
</tr>
<tr>
<td>4. Enthusiasm</td>
<td>0.93</td>
<td>3.76</td>
<td>1.18</td>
<td>–0.10</td>
<td>–0.57</td>
<td>–0.51 0.82 –0.47</td>
</tr>
</tbody>
</table>

* Two weeks after the first questionnaire completion (Time 1).
α – Cronbach’s α; M – mean; SD – standard deviation; SKE – skewness; K – kurtosis.
* All correlations are statistically significant at least at the level of p < 0.01.

varied between 0.92 and 0.78, for comfort varied between 0.9 and 0.78, for depression varied between 0.91 and 0.82, and for enthusiasm varied between 0.92 and 0.9, all being statistically significant. The values of AVE ranged from 0.68 for comfort to 0.82 for enthusiasm. The values of CR ranged from 0.86 for comfort to 0.93 for enthusiasm.

To check if each scale captures distinct constructs, the square root of AVE for the scale is compared with its correlations with other scales [57]. If it is higher than correlations the discriminant validity is achieved. The square roots of AVE of the scales are 0.83 for comfort, 0.85 for anxiety, 0.87 for depression, and 0.9 for enthusiasm. They all exceed values of correlations between scales (Table 4).

**Criterion and predictive validity**

To test the criterion validity of the Polish version of Warr’s job-related affective well-being measure, the relationships between its scales and the criterion variables, mood and turnover intention, have been evaluated (Table 5).
Summing up, the results have confirmed all hypothesis concerning relationships between job-related affect and mood as well as turnover intentions. This reveals that the Polish version of the Warr measure demonstrates criterion and predictive validity capturing 4 affective dimensions in accordance with underlying theory.

**Discussion**

The aim of the current study was to test the factorial structure of job-related affect in a Polish sample and to present the Polish adaptation of Warr’s job-related affective well-being measure. Firstly, we checked psychometric properties of each item of the measure. Next, we compared the goodness of fit of 6 alternative factorial models and chose the best fitting one for further analyses of reliability and validity of the measure.

Examination of alternative models confirm that Warr’s job-related affective well-being measure captures 4 dimensions measuring enthusiasm, comfort, anxiety, and depression as affective states different in valence and in activation level (Figure 1). Moreover, there are also other studies that confirm the 4-factor model as the best fitting to the data in samples from Poland, Spain and the Netherlands [58]. Therefore, the
factorial structure of the Polish version of the measure is firmly established. Moreover, the results of this study reveal that 4 scales of the Polish adaptation of the instrument achieve acceptable levels of reliability and validity. Values of Cronbach’s α indicate good to excellent reliability of scales. Test-retest correlations indicate acceptable levels of temporal stability over a 2 week period [7]. This shows that affect at work is relatively stable, which may be treated as a bit unexpected result, but it is consistent with the results of other studies [59]. Correlations between scales of job-related affective well-being measure and turnover intention as well as mood show a pattern of results in line with theoretical expectations [20,40,44,45]. Higher levels of enthusiasm and comfort and lower levels of anxiety and depression may be predicted by positive mood. Conversely, negative mood fosters the growth of enthusiasm and comfort and the decrease of anxiety and depression. In addition, enthusiasm and comfort lower turnover intention and anxiety and depression boosts it in the future.

Limitations and future directions

All in all our study shows that the Polish adaptation of the job-related affective well-being measure has resulted in a psychometrically valid tool. Additional research, however, is welcome to confirm discriminant validity of the scales. For instance, relations between the job-related affect scales and other scales assessing affect at work (e.g., JAWS) [5] may be analyzed in future studies. There are results confirming partial metric measurement invariance of the scales across samples of employees from 3 European countries (including Poland) and full scalar invariance across gender [58]. It is open, however, for further studies to attest measurement invariance of the instrument across other samples, for example across age groups. This study has attested to the psychometric properties of the well-known 12-item Warr [9] measure, which has many national adaptations [4,7,8]. There is, however, also another 16-item instrument published recently by Warr et al. [60]. This is an extended measure of job-related affect that awaits further research and national adaptations which confirm its validity.

CONCLUSIONS

To sum up, Warr’s [9] job-related affective well-being measure is an appropriate tool for measuring affect experienced at work as a multi-dimensional phenomenon. Contrary to other popular scales which capture context-free positive and negative affect (e.g., PANAS), this measure allows researchers to assess 4 dimensions of affect in a work context. This instrument may be useful for future studies in samples of Polish-speaking employees of various professions, including self-employed, as this has been validated in other studies [17,58]. Availability of the Polish adaptation of this measure may contribute to further development of research on affect in a work context, which is a dynamically growing area of investigation [2–4].

REFERENCES


